Politics, government health expenditure and infant mortality:

Does political party orientation matter?

Constantinos Alexiou<sup>a</sup> and Emmanouil Trachanas<sup>b,\*</sup>

<sup>a</sup> Cranfield School of Management, Cranfield University, Cranfield MK43 0AL, UK. E-mail:

constantinos.alexiou@cranfield.ac.uk.

<sup>b</sup>Oxford Brookes Business School, Oxford Brookes University, Oxford OX3 0BP, UK. E-mail:

etrachanas@brookes.ac.uk. \* Corresponding author.

International Journal of Social Economics

DOI: 10.1108/IJSE-04-2021-0221

Abstract

Purpose: Motivated by the scant available evidence we explore the relationship between

government political party orientation and infant mortality.

Design/methodology/approach: We apply a panel quantile methodology to a dataset that

consists of 15 countries of the G20 group over the period 2000-2018. We control for

heterogeneous parameters across countries and quantiles and obtain estimates across the

different points of the conditional distribution of the dependent variable.

Findings: Our findings support the hypothesis that political party orientation has a significant

effect on a population health indicator such as infant mortality. The analysis suggests that, to a

great extent, left-wing government parties contribute to better health outcomes - when

compared to right and centre political parties - both individually as well as interacted with

government health expenditure. Moreover, the impact of redistributing policies appears to be

of a paramount importance in alleviating infant mortality while more education and lower

unemployment can also contribute to better health outcomes.

**Originality:** We explore the relationship between the nature of government political party

orientation (i.e. right, centre and left) and infant mortality whilst at the same time gauging the

mediating effect of party orientation via government health expenditure on infant mortality.

Additional aspects of the impact of other control variables such as, income inequality,

unemployment, and education on infant mortality are also investigated.

**Keywords:** health expenditure; infant mortality; panel data

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#### Introduction

The literature on the determinants of health outcomes is well developed while there is growing evidence about the impact of social and cultural factors on population health. However, only scant evidence exists about the impact of government political party orientation on health outcomes while the distribution and exercise of power in health policy and systems has received little attention (Shiffman, 2014; Lee, 2015). Although several empirical studies have explored the potential channels through which the nature of political regimes (democratic, autocratic, etc.) affects economic development (e.g., Przeworski *et al.*, 2000; Acemoglu and Robinson, 2006) or health outcomes (e.g. Okada, 2018), studies that investigate the impact of government political party orientation on population health have fallen behind. Given the nature as well as the composition of current democratic societies one would have expected more emphasis to be placed on the policy implications of democratically elected political parties on health outcomes. Rudolf Virchow – one of the most influential figures in social medicine – eloquently summarized that "Medicine is a social science and politics nothing but medicine on a grand scale".

One possible explanation for the limited literature on the impact of politics on health outcomes might relate to the fact that the public health field is mainly dominated by medical researchers who consider politics to be controversial or falling outside their remit. An alternative explanation might also reside in the way most public health research is funded. For instance, Navarro (2008) argues that within the European region most of the funding comes from public sources that are closely related to political parties. It would therefore be unlikely that an incumbent administration of one political party approves research evidence that suggests that parties of the opposite orientation are more effective in improving population health.

Arguably, the exercise of power might indeed be a feature of actors or institutions that have the ability to shape global health policy (Brown *et al.*, 2006), but at a national level political parties and their concomitant ideologies can be of paramount importance in influencing the equity and universality of public policy (Navarro *et al.*, 2006; Mackenbach, 2014). It is in this sense that the World Health Organization's Director-General suggested that universal health coverage (UHC) is "a political choice" (UN News 2019; WHO 2019).

Politics is a ubiquitous feature of modern societies whilst political views and economic systems differ markedly. The traditional ideological dichotomy between left-wing and right-wing party orientation reflects the degree of government intervention in a market system to either secure social goals or provide more economic freedom and minimal state intervention.

In this context, existing social and economic inequalities are believed to be strong predictors of health inequalities whilst increased income inequality is found to adversely affect both health and social outcomes - such as infant mortality, life expectancy, obesity, mental health, trust, level of education and social mobility (Willkison and Pickett, 2010). There is also evidence that political orientation in conjunction with personal and vested interests can have a substantial impact on health policy (Ferrie, 2015; Reijneveld, 2017).

From a policy perspective, the proponents of the view that institutions are closely bound up with the evolution of welfare systems suggest that direct democracies, and federalism, tend to promote a more liberal type of welfare state by vetoing decisions and by steering the central government's direction to the expansion of social expenditure (Immergut, 1992; Wagschal, 1997; Crepaz, 1998; Lijphart, 2012; Castles, 2000; Tsebelis, 2002). In view of the above and the very few studies in the area, such as the one by Korpi (2000) who explored how political traditions and the power relations reflected by class as well as gender affect the nature of the welfare state and the well-being of the population, we feel that there is a scope and a pressing need to undertake more studies exploring the impact of politics on health outcomes.

In this direction we attempt to contribute to the existing literature in a number of ways. First, by applying a panel quantile methodology to a dataset that consists of 15 countries of the G20 group over the period 2000-2018, we control for heterogeneous parameters across countries and quantiles and obtain estimates across the different points of the conditional distribution of the health indicator variable. Second, to the best of our knowledge this is the first study that explores the relationship between the nature of government political party orientation (i.e. right, centre and left) and infant mortality whilst at the same time gauges the mediating effect of party orientation via government health expenditure on infant mortality. Additional aspects of the impact of other control variables such as, income inequality, unemployment, and education on infant mortality are also investigated.

Our findings support the hypothesis that political party orientation affects indicators of population health such as infant mortality. Our analysis makes an empirical link between politics and policy by showing that to a great extent left-wing parties via health expenditure outlays produce better health outcomes compared to right and centre oriented political parties. In addition, the impact of redistributing policies (as reflected by the highly significant Gini coefficient) appears to be of a paramount importance in alleviating infant mortality whilst both education and unemployment incrementally reinforce the hypothesis that more education and lower unemployment contribute to better health outcomes.

The remainder of this paper is organized as follows. The second section presents a brief overview of the related literature whilst the third section describes the data and the estimation methodology adopted in this study. The fourth section discusses the estimation results and the final section provides some concluding remarks.

#### Relevant literature

The preponderance of the existing health policy and systems analyses are mainly descriptive (Adam *et al.*, 2012), lacking theoretical foundations that can generate durable intellectual frameworks to analyse health problems in different contexts (Kleinman, 2010), making their suggestions "incoherent and difficult to integrate into foreign policy strategies" (Youde, 2005, p. 205). Gilson and Raphaely (2008) reviewed the policy analysis literature and found that a vigorous assessment of politics and power was rarely studied whilst Storeng and Mishra (2014) observed that health systems strengthening studies focused on technical and managerial issues of health service delivery, whilst no attention was paid to the politics and social relations that affect health systems.

According to the hypothesis of partisan influence on public policy one key factor accounting for significant variations in policy outcomes in constitutional democracies is the party composition of government (see, Hibbs, 1977; Castles, 1982; Cameron, 1984; Hicks and Swank, 1992; Schmidt, 1996; Okada, 2018). In this context, right-wing political parties are associated with significantly lower welfare expenditure whilst left-wing parties tend to encourage higher welfare spending.

A report from the World Health Organization (2008) attributed health inequities to a toxic environment that nurtures poor social policies, unfair economic arrangements, and bad politics. Even though the report conceptualised and collated qualitative evidence on a wide range of social forces and casual processes that can potentially affect health inequity, aspects on competing political priorities, contested political ideologies, or other factors encountered when moving forward health policy agendas, were rather ambiguous (Lee, 2010). In a recent report, the World Health Organizations (2018) used the same rhetoric (i.e., that health inequities are attributed to factors including education, employment status, income level, gender and ethnicity) without however making an effort to provide clear policy insights relating to specific political or ideological persuasion.

Ottersen *et al.* (2014) defined the "global political determinants of health as the norms, policies, and practices that arise from global interactions among entities (states, transnational corporations, and civil society organisations, among others) with different interests and degrees

of power" (p.630). The existing health reports in the area have to a great extent failed to "consider the need to better understand empirically how power is expressed in global health governance" (Marten, *et al.*, p. 2207).

Developing interdisciplinary research that can systematically explore issues relating to inequity and political power or generate a decent dialogue between global health researchers requires that humanities and social sciences are combined (Ooms, 2014). Despite the fact that over the past few years there has been a growing interest in social-scientific scholarship on questions of power and politics in global health (Hansen *et al.*, 2013; Storeng and Mishra, 2014; Gomez, 2016; Parker and Garcia, 2019), there are voices that are calling for political science "to take global health policy issues more seriously, while the global public health community needs to do the same for political science, specifically on issues of institutional design and governance" (Gomez, 2016, p. 4).

Despite the fact that extensive theoretical research has been conducted on how political traditions affect health policy, hardly any empirical studies have emerged that explore the impact of political party orientation on health outcomes. At a policy level, the evidence is readily available and as Navarro (2006) argues, parties promoting social policies and a fairer distribution of income have been more successful in improving health population compared to parties whose policies are devoid of such a commitment. On the empirical front, single-country evidence on the health effects of political party orientation (see, McGuire 2010; Scott-Samuel et al., 2014; Hiam et al., 2017; Watkins et al., 2017) can offer useful insights, but internationally comparative studies can provide a more holistic analysis by transcending the specificities of individual countries. Muntaner et al. (2011) explored four key political features i.e., democracy, welfare state, left-of-centre political tradition and globalisation, and found that a) a democracy may be health-promoting due to the inherent ability of the electorate to use their voting power when faced with unpopular policies, b) the left-of-centre political tradition is directly associated with a welfare state which promotes better health outcomes through policies that reduce social, economic and health inequalities (Rowlingson, 2011), and c) globalisation, as this is reflected by the emergence of free-trade and neoliberal policies, is advocated by right-wing political entities the policies of which exacerbate population health outcomes. In the same spirit, Barnish and Turner (2017) concluded that a generous welfare state, left-of-centre democratic political tradition, and democracy, appear to be positively affecting population health whilst the evidence on globalisation may be less conclusive.

Moreover, another study using data for various OECD economies governed mainly by social democratic parties (Sweden, Norway, Denmark, Finland and Austria) suggested that

their commitment to redistributive policies as measured by the average Gini coefficient over the last 10 years of the study period (which was 0.22) (LIS, 2006) provided universal health care coverage, and social benefits to all citizens (ranging from 30% of GDP for social public expenditure to 7.2% of GDP for the average public health care expenditure over the last 10 years of the study period). More recent studies suggest that Nordic countries rank high among OECD countries in terms of per capita expenditure on health (averaging 10.3% of GDP) whereas Austrian health expenditure in 2016 was 10.4% of GDP respectively (OECD 2017; Morgan et al. 2017). In this context, Raphael and Bryant (2004) argue that social democratic parties nurture policies that support women's health and wellbeing, paid maternity leave, early child education, childcare, low crime, participation of women in government, and home care services. On the other hand, those economies (i.e., Italy, Netherlands, Germany, Belgium and France) that have been less committed to redistributive policies (average Gini coefficient of 0.30) than social democrats, have been mainly governed by conservative parties or Christian democratic parties for most of the period from 1950 to 2000. Indicatively, the average public social expenditure and average public health care expenditure were 28% and 6.4% of GDP respectively (Bambra, 2005). Recently the average public heath expenditure in these countries has increased further registering 8.5% of GDP (Morgan et al, 2017).

Furthermore, countries mainly run by liberal or conservative parties of a liberal persuasion, (such as the UK, Ireland, Canada and the USA) appear to be adopting policies that are redistributive in nature (the group average Gini coefficient was 0.32) whilst public social expenditures are much lower than in the countries governed by social democratic and Christian democratic parties. In particular, average public social expenditure was 24% of GDP, whereas the average public health care expenditure amounted to 5.8% of GDP (Esping-Andersen, 1990). The picture however concerning the respective economies appears to have changed dramatically as more recent data suggests that the average health expenditure has doubled, reaching 11.6% of GDP (OECD 2017).

Countries governed for a considerable period of time by conservative dictatorships (such as Spain, Portugal Greece) during the period of the study (1950-2000) were characterized by unequal income distribution, underdeveloped welfare state and poor public services (the average Gini coefficient was 0.42). In these group of countries public social expenditures were very low (the average social expenditure at the end of each dictatorship, sometime in the 1970s, was only 14% of GDP, and the average public health care expenditure was only 4.8% of GDP). It should be stressed however that in the years that followed, democracy was established in these countries and their average public social expenditure and average public health care

expenditure were close to those of the liberal tradition (see, Navarro, 1989, 2002) while more recently health expenditure has amounted to 8.7% of GDP (Morgan *et al.* 2017).

### Data and methodology

We investigate the effects of government health expenditure and government party political orientation on health outcomes by estimating an unbalanced panel with annual data for the period 2000-2018. We focus on 15 countries of the G20 group<sup>1</sup> (Argentina, Australia, Brazil, Canada, France, Germany, India, Italy, Japan, Korea Republic, Mexico, South Africa, Spain, United Kingdom, USA).

Our baseline regression model is:

health<sub>it</sub> =  $\alpha_i + \beta_1 govhealthex_{it} + \beta_2 inequality_{it} + \beta_3 unemployment_{it} + \beta_4 education_{it} + \beta_5 \mathbf{X}_{it} + \varepsilon_{it}$  (1) where health denotes health outcomes captured by infant mortality, govehealthex denotes government health expenditure as a percentage of GDP, inequality denotes a Gini index of income inequality in equivalized household disposable income developed by Solt (2020), unemployment is the unemployment rate, and education is secondary school enrolment.

Equation (1) also includes a vector **X** that consists of dummy variables that capture government party orientation with respect to economic policy which are constructed using data from the Database of Political Institutions 2020 (Cruz *et al.*, 2021). Specifically, *right* is a dummy variable that takes the value 1 if the government party is defined as right (conservative, Christian democratic, right-wing) and 0 otherwise (Model 1); *centre* takes the value 1 if the government party position can best be described as Centrist and 0 otherwise (Model 3); and *left* takes the value 1 if the government party is left (communist, socialist, social democratic, left-wing) and 0 otherwise (Model 5). Furthermore, by interacting the types of political party in power with government health expenditure, we explore the potential effect that may be manifested through the impact of the interacting terms (Models 2, 4, 6). Table A1 in the Appendix presents detailed descriptions of all variables and their sources, while Tables A2 and A3 provide descriptive statistics and the correlation matrix.

Traditional regression techniques (e.g. OLS) measure differences in the dependent variable between populations at the mean and in this setting the fundamental assumption made is that the regression coefficients are constant across the population. However, in some cases we are interested to investigate if the independent variables have varying effects at different points in the conditional distribution of the dependent variable, rather than only at the mean.

<sup>&</sup>lt;sup>1</sup> Our effort to use data for all the G20 countries and a longer time period was hampered by data availability.

This is important in our context since it is reasonable to expect that the effect of government health expenditure may differ across quantiles of the conditional distribution of the health outcome variable. To this end, we employ a quantile panel estimator with nonadditive fixed effects which produces estimates at different points of the conditional distribution of the dependent variable which are considered to be more robust to nonnormal errors and outliers.

Many panel quantile estimators include additive fixed effects in the quantile function and provide estimates about the distribution of  $(Y_{it} - \mu_i | Z_{it})$ , where  $Y_{it}$  is the outcome variable,  $Z_{it}$  are treatment variables (exogenous or endogenous), and  $\mu_i$  denotes the fixed effects. However, observations at the top of  $(Y_{it} - \mu_i)$ , may actually be at the bottom of the  $Y_{it}$  distribution and, subsequently, additive fixed effect models can provide information about the outcome relative to fixed effect distribution rather than the effects of the treatment variables on the outcome distribution. In this paper we employ Powell's (2016) panel data quantile estimator with nonadditive fixed effects that provides estimates for the distribution of  $(Y_{it}|Z_{it})$ , while allowing for individual level heterogeneity and maintaining the non-separable disturbance term typically used in quantile estimation. The quantile panel estimator with nonadditive fixed effects produces estimates which are consistent for small T and can be interpreted in the same manner as cross-sectional quantile estimates, i.e. the impact of the explanatory variables on a quantile of the outcome distribution.

### Results and discussion

We use infant mortality as a proxy for health outcomes and as implied by the quantile methodology adopted, we estimate alternative models at different point of the distribution (0.10, 0.25, 0.50, 0.75, 0.90). In Models 1, 3, and 5 (Tables 1, 2, 3) the four control variables are common in all models whilst the party orientation dummies are inserted one at a time. In Models 2, 4, and 6 we use the same rationale as previously but this time the party dummy variables are interacted with government health expenditure in order to capture the mediating effect of party orientation through government health expenditure on infant mortality.

### [Tables 1, 2, 3 about here]

The results indicate that a first key finding is the paramount and highly significant role that government health expenditure, income inequality, and party orientation play in affecting infant mortality. More specifically, the empirical results show that the relationship between government health expenditure (*govhealthex*) and infant mortality is robust and statistically significant with varying degrees of negative impact across the five quantiles for each of the

models estimated. Our results suggest that at least for the set of counties in our sample, effective allocation of government health expenditure will indeed reduce infant mortality a finding in line with Rana *et al.* (2018), Nicholas *et al.* (2016), Rahman *et al.* (2018) and Owusu *et al.* (2019) but inconsistent with the study of Akinlo and Sulola (2019).

Income inequality (*inequality*) is found to be positive and highly significant across all estimated models and different points of the distribution which is in line with Kawachi and Kennedy (1999) who argue that redistributive economic policies have a substantial impact on social justice and better population health. The proposition that income distribution plays a key role in determining population's health, is also explored by Mellor and Milyo (2002), Osler *et al.* (2002), Shibuya *et al.* (2002); Deaton and Lubotsky (2003), Lynch *et al.* (2004), Leigh *et al.* (2012), Pickett and Wilkinson (2015), Neumayer and Plumper (2016), amongst others. It is therefore envisaged that the countries where policies distribute income more equitably will be boasting a heathier population.

Unemployment is found to have a positive effect across the majority of the estimated models which is in line with Adofu and Salami (2019). It is worth noting here that the broader literature on unemployment proxied by economic downturns and child health has produced mixed evidence (see for instance, Dehejia and Lleras-Muney 2004 and Eiríksdóttir *et al.* 2013. Furthermore, education has a significantly negative effect however this finding is relatively difficult to be compared against existing studies on the relationship between education and infant mortality as the majority, if not all of these, use survey data that consists of various socioeconomic factors such as mother's education, mother's work status, husband's occupation, husband's education, type of place of residence etc. The level of mother's education in a number of this these studies is particularly strongly associated with infant mortality whilst in some others it is the husband's education that is more important (Hobcraft *et al.*, 1984; Arntzen and Andersen, 2004; Gakidou *et al.*, 2010; Andriano *et al.*, 2019).

The party orientation dummies exhibit interesting patterns when they are examined individually or as interaction terms with government health expenditure. In particular, the party right dummy (*right*) when considered individually appears to be initially negatively related to infant mortality at the 0.10 and 0.25 quantiles respectively, then turns positive and significant at the 0.5 quantile before it becomes again negative and significant at the 0.75 quantile (Table 1, Model 1). When the party interactions are considered, the coefficient of the party right interaction (*right*×*govhealthex*) is negative and significant at the lowest quantile 0.10 which then changes to positive at quantiles 0.50 and 0.75 (Table 1, Model 2). The party centre dummy (*centre*) is found to be positive and significant at the first two quantiles 0.10 and 0.25 before it

turns negative at quantiles 0.50 and 0.90 (Table 2, Model 3). The estimates for the interaction term (*centre*×*govhealthex*) are more fragmented since it is found to be negative and highly significant only at the 0.50 quantile (Table 2, Model 4). The party left dummy (*left*) is found to be positive and significant at the lowest quantile 0.10 before it becomes negative and highly significant at the following three quantiles 0.25, 0.50, and 0.75 (Table 3, Model 5) while the results of the interaction term (*left*×*govhealthex*) are mostly consistent; positive and significant at the lowest quantile 0.10 and change to negative and significant at the following quantiles 0.25, 0.50, and 0.75 (Table 3, Model 6).

Our results to a great extent confirm the view that left-wing governments promote, through health expenditure provision, better health outcomes than under right-wing governments. This is evident in the highly significant coefficients of the individual left party dummy as well as the interaction term in the respective quantiles (0.25, 0.50, 0.75). In this context, infant mortality may be influenced by leftist political orientation as they tend to support policies that favour the weakest and disadvantaged of the populations (Subramanian *et al.*, 2009). The fact however that at the lowest quantile 0.10 the interaction term was found to positively affect infant mortality might suggest that certain segments of the society will not benefit from higher spending. Even though it might not directly follow from our results it might be inferred that people at the lowest end of the distribution are unable to strike a balance between individual lifestyle choices versus wider structural factors that shapes their health outcome, hence, being unable to harness the benefits of higher health expenditure.

Overall, the policy implications of our study are of great significance as it raises awareness of the fact that *politics do matter* and should therefore assume prominence in public health research. It is impossible to understand health outcomes without acknowledging the ideological differences between governments. Our study clearly suggests that state administrations, such as left-wing governments, that favour redistributional policies (through more progressive tax systems and universal welfare cultures) are more likely to be associated with improved population health outcomes such as lower rates of infant mortality. Policy makers need to understand that the assumption that democratic systems serve as an automatic remedy of ailments, is mostly a misconception. Infant mortality can only be reduced by the right mix of government policies. Well-designed policies that aim at boosting social and health expenditure should be envisaged as the precursors of improved health outcomes.

#### **Conclusions**

Arguably, government healthcare expenditure is one of the most important components of total government expenditure in advanced economies and potentially one of the most controversial topics in the political discourse. The notion that partisan ideology is associated with changes in health expenditure in the OECD (Herwartz and Theilen, 2014) provided the impetus for this study to assess its mediating effect on health outcomes.

Research on the social determinants of health (Braveman *et al.*, 2011) has highlighted how environmental and societal factors can shape health outcomes. Insights into the institutional aspects of racism and discrimination and their concomitant impact on health disparities have also been provided (Bailey *et al.*, 2017). In this study we have gone one step further by assessing the political determinants of health as an important, but currently underresearched, determinant of health. We have explored the relationship between the government political party orientation and health outcomes (as proxied by infant mortality) and gauged the mediating effect of party policy via government health outlays on infant mortality. Additional aspects potentially significant for health outcomes such as income inequality, unemployment, and education were also investigated.

Our findings support the hypothesis that political party orientation affects indicators of population health outcomes such as infant mortality. We find that left-wing parties, when compared to right and centre oriented political parties, contribute to better health outcomes. In addition, the impact of redistributing policies (as reflected by the highly significant Gini coefficient) appears to be of a paramount importance in alleviating infant mortality whilst both education and unemployment incrementally reinforce the premise that more education and lower unemployment can lead to better health outcomes.

Setting priorities for health spending has been at the forefront of research for decades and health policy researchers have been trying to come up with the proper formula that could be used to effectively plan health care expenditure policies. Despite the existing limitations that relate to the measures of burden - that emphasize mortality or health expenditure - it is apparent that no formulas or broad principles can be used in lieu of a political process when determining priorities in health care expenditure. When it comes to which party ideology favours a more equitable and healthy society, left-wing ideology appears to dominate the theoretical we well as empirical discourse of academic analysis. The question, however, is not to get entangled into party politics that most of the time are dominated by vested interests but rather to devise a political process that is not dominated by those with greater wealth and position of power. We

should find a way to a process that is fair and transparent given the history of health policy across the world.

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### **Tables**

Table 1. Panel quantile results for right government party policy orientation

| Model 1      |           |           |           |           |           |  |
|--------------|-----------|-----------|-----------|-----------|-----------|--|
| Quantile     | 0.1       | 0.25      | 0.5       | 0.75      | 0.9       |  |
| govhealthex  | -0.297*** | -0.444*** | -0.474**  | -0.356*** | -2.839*** |  |
|              | (0.078)   | (0.032)   | (0.238)   | (0.089)   | (0.778)   |  |
| inequality   | 0.844***  | 0.753***  | 0.708***  | 0.582***  | 0.850***  |  |
|              | (0.045)   | (0.009)   | (0.027)   | (0.046)   | (0.092)   |  |
| unemployment | 0.002     | 0.080***  | 0.197***  | 0.691***  | 0.098***  |  |
|              | (0.031)   | (0.006)   | (0.052)   | (0.044)   | (0.019)   |  |
| education    | -0.069*** | -0.085*** | -0.076*** | -0.255*** | 0.073     |  |
|              | (0.021)   | (0.005)   | (0.022)   | (0.018)   | (0.118)   |  |
| right        | -0.611**  | -0.675*** | 0.674**   | -0.734*** | 0.905     |  |

| V | Ιo | d | el | 2 |
|---|----|---|----|---|
|---|----|---|----|---|

(0.287)

(0.137)

(0.657)

(0.146)

(0.242)

| Quantile          | 0.1       | 0.25      | 0.5       | 0.75      | 0.9       |
|-------------------|-----------|-----------|-----------|-----------|-----------|
| govhealthex       | -0.481*** | -0.336*** | -0.312*** | -0.234*** | -2.685*** |
|                   | (0.042)   | (0.106)   | (0.068)   | (0.049)   | (0.354)   |
| inequality        | 0.809***  | 0.700***  | 0.676***  | 0.800***  | 0.721***  |
|                   | (0.032)   | (0.034)   | (0.029)   | (0.009)   | (0.099)   |
| unemployment      | -0.011    | -0.072    | 0.160***  | 0.541***  | 0.326**   |
|                   | (0.059)   | (0.095)   | (0.025)   | (0.024)   | (0.141)   |
| education         | -0.075*** | -0.050    | -0.092*** | -0.134*** | -0.031    |
|                   | (0.011)   | (0.048)   | (0.006)   | (0.004)   | (0.044)   |
| right×govhealthex | -0.135*** | -0.020    | 0.100***  | 0.052***  | 0.100     |
|                   | (0.021)   | (0.048)   | (0.019)   | (0.007)   | (0.099)   |

Notes: 206 observations in total. Standard errors in parentheses. The adaptive Markov Chain Monte Carlo optimization procedure is employed with 1000 draws. \*\*\*, \*\* and \* denote statistical significance at the 1, 5 and 10% level, respectively.

Table 2. Panel quantile results for centre government party policy orientation

## Model 3

| Quantile     | 0.1       | 0.25      | 0.5       | 0.75      | 0.9        |
|--------------|-----------|-----------|-----------|-----------|------------|
| govhealthex  | -0.519*** | -0.381*** | -1.178*   | -0.531*** | -2.493***  |
|              | (0.067)   | (0.033)   | (0.614)   | (0.073)   | (0.190)    |
| inequality   | 0.781***  | 0.819***  | 0.608***  | 0.755***  | 0.765***   |
|              | (0.033)   | (0.007)   | (0.074)   | (0.013)   | (0.053)    |
| unemployment | -0.010    | -0.003    | 0.392*    | 0.523***  | 0.053      |
|              | (0.066)   | (0.011)   | (0.225)   | (0.009)   | (0.047)    |
| education    | -0.087*** | -0.050*** | -0.097*** | -0.116*** | -0.151***  |
|              | (0.015)   | (0.005)   | (0.018)   | (0.003)   | (0.010)    |
| centre       | 0.980***  | 0.854***  | -1.337**  | 0.756     | -13.386*** |
|              | (0.217)   | (0.224)   | (0.663)   | (0.759)   | (1.164)    |

### Model 4

| Quantile           | 0.1       | 0.25      | 0.5       | 0.75     | 0.9       |
|--------------------|-----------|-----------|-----------|----------|-----------|
| govhealthex        | -0.499*** | -0.553*** | -0.419*** | -1.503** | -2.465*** |
|                    | (0.147)   | (0.040)   | (0.155)   | (0.604)  | (0.646)   |
| inequality         | 0.659***  | 0.799***  | 0.686***  | 1.163*** | 0.796***  |
| •                  | (0.089)   | (0.017)   | (0.038)   | (0.199)  | (0.159)   |
| unemployment       | -0.020    | -0.004    | 0.124*    | 0.110    | 0.280     |
|                    | (0.054)   | (0.043)   | (0.065)   | (0.180)  | (0.193)   |
| education          | -0.083*** | -0.018    | -0.104*** | 0.059    | 0.012     |
|                    | (0.021)   | (0.027)   | (0.007)   | (0.088)  | (0.076)   |
| cenrte×govhealthex | 0.194     | -0.115*   | -0.369*** | -0.270   | 0.495     |
|                    | (0.148)   | (0.059)   | (0.040)   | (0.948)  | (1.852)   |

Notes: See notes Table 1

Table 3. Panel quantile results for left government party policy orientation

## Model 5

| Quantile     | 0.1       | 0.25      | 0.5       | 0.75      | 0.9       |
|--------------|-----------|-----------|-----------|-----------|-----------|
| govhealthex  | -0.453*** | -0.019    | -0.481*** | -0.119*** | -2.299*** |
|              | (0.095)   | (0.080)   | (0.153)   | (0.030)   | (0.816)   |
| inequality   | 0.869***  | 0.797***  | 0.774***  | 0.826***  | 0.828***  |
|              | (0.043)   | (0.021)   | (0.064)   | (0.005)   | (0.148)   |
| unemployment | 0.011     | -0.039*** | 0.372***  | 0.424***  | 0.165     |
|              | (0.039)   | (0.011)   | (0.128)   | (0.010)   | (0.194)   |
| education    | -0.078*** | -0.089*** | -0.025    | -0.131*** | -0.012    |
|              | (0.015)   | (0.012)   | (0.038)   | (0.003)   | (0.074)   |
| left         | 0.807***  | -0.766*** | -0.960**  | -0.083**  | -0.320    |
| -            | (0.231)   | (0.216)   | (0.455)   | (0.039)   | (0.738)   |

## Model 6

| Quantile         | 0.1       | 0.25     | 0.5       | 0.75      | 0.9       |
|------------------|-----------|----------|-----------|-----------|-----------|
| govhealthex      | -0.603*** | 0.062    | -0.123*** | -0.155*** | -1.687*** |
|                  | (0.033)   | (0.115)  | (0.037)   | (0.009)   | (0.325)   |
| inequality       | 0.807***  | 0.923*** | 0.697***  | 0.830***  | 1.229***  |
|                  | (0.013)   | (0.050)  | (0.011)   | (0.002)   | (0.221)   |
| unemployment     | -0.020    | -0.102*  | 0.155***  | 0.459***  | -0.073    |
|                  | (0.032)   | (0.057)  | (0.018)   | (0.008)   | (0.167)   |
| education        | -0.089*** | -0.059** | -0.093*** | -0.135*** | 0.106     |
|                  | (0.003)   | (0.025)  | (0.004)   | (0.002)   | (0.122)   |
| left×govhealthex | 0.100***  | -0.077** | -0.081*** | -0.029*** | 0.031     |
|                  | (0.015)   | (0.035)  | (0.010)   | (0.008)   | (0.100)   |

Notes: See notes Table 1

# Appendix

Table A1. Description of variables

| Variable     | Definition  | Source   |
|--------------|---|--|
| health       | Infant mortality rate is the number of infants dying before reaching one year of age per 1,000 live births.   | World Development Indicators   |
| govhealthex  | General government health expenditure from domestic sources (% of GDP).   | World Development Indicators   |
| inequality   | Gini index of inequality based on household disposable (post-tax, post-transfer) income.  | Standardized World Income<br>Inequality Database 9.0 (Solt,<br>2020)                                       |
| unemployment | Unemployment (% of total labour force).   | World Development Indicators   |
| education    | School enrolment, secondary (% gross).  | World Development Indicators   |
| right        | Government party economic policy orientation dummy variable that takes the value 1 if the government party is right (conservative, Christian democratic, right-wing) and 0 otherwise.             | Constructed by the authors using data from the Database of Political Institutions 2020 (Cruz et al., 2021) |
| centre       | Government party economic policy orientation dummy variable that takes the value 1 if the government party can be described as Centrist and 0 otherwise.  | Constructed by the authors using data from the Database of Political Institutions 2020 (Cruz et al., 2021) |
| left         | Government party economic policy orientation<br>dummy variable that takes the value 1 if the<br>government party is left (communist, socialist,<br>social democratic, left-wing) and 0 otherwise. | Constructed by the authors using data from the Database of Political Institutions 2020 (Cruz et al., 2021) |

**Table A2. Summary statistics** 

| -            | Table 182. Summary statistics |       |          |       |        |  |  |  |
|--------------|-------------------------------|-------|----------|-------|--------|--|--|--|
| Variable     | Obs.                          | Mean  | St. Dev. | Min   | Max    |  |  |  |
| health       | 285                           | 11.71 | 13.91    | 1.80  | 66.70  |  |  |  |
| govhealthex  | 285                           | 5.52  | 2.27     | 0.71  | 9.21   |  |  |  |
| inequality   | 275                           | 37.79 | 9.43     | 26.20 | 63.50  |  |  |  |
| unemployment | 267                           | 8.75  | 6.43     | 2.40  | 33.29  |  |  |  |
| education    | 233                           | 98.59 | 16.0     | 44.87 | 157.20 |  |  |  |
| right        | 268                           | 0.478 | 0.50     | 0     | 1      |  |  |  |
| centre       | 268                           | 0.082 | 0.27     | 0     | 1      |  |  |  |
| left         | 268                           | 0.440 | 0.49     | 0     | 1      |  |  |  |
|              |                               |       |          |       |        |  |  |  |

**Table A3. Correlation matrix** 

| Variable     | health | govhealthex | inequality | unemployment | education |
|--------------|--------|-------------|------------|--------------|-----------|
| health       | 1.00   |             |            |              |           |
| govhealthex  | -0.66  | 1.00        |            |              |           |
| inequality   | 0.88   | -0.63       | 1.00       |              |           |
| unemployment | 0.49   | -0.12       | 0.57       | 1.00         |           |
| education    | -0.59  | 0.51        | -0.43      | 0.09         | 1.00      |